

Lecture 2: (Chap 1, cont.)

In this chapter the fact that Brownian motion is a continuous Gaussian process will play an important role.

Definition:

An \mathbb{R}^d -valued process, $X_t, t \in T$, (T is some arbitrary non-empty set), defined on (Ω, \mathcal{A}, P) is a centered Gaussian process, if for any $n \geq 1$, $t_1, \dots, t_n \in T$, $\lambda_1, \dots, \lambda_n \in \mathbb{R}^d$, $\sum_{i=1}^n \lambda_i \cdot X_{t_i}$ is a real-valued centered Gaussian variable (possibly $\equiv 0$).
↑ scalar product when $d \geq 2$

The $d \times d$ -matrix valued function on T^2 :

$$(1.12) \quad \Gamma(u, v) = E[X_u \cdot X_v] = (E[X_u^i X_v^j])_{1 \leq i, j \leq d}, \quad u, v \in T$$

(note that $\Gamma(v, u) = {}^t \Gamma(u, v)$),

is the covariance function of the process,

(note that $E[X_t] = 0$, for each $t \in T$).

Lemma: $(X_t)_{t \in T}$ a centered Gaussian process with covariance function Γ)

The function $\Gamma(u, v), u, v \in T$, completely determines all
 (1.13) finite distributions X_{t_1}, \dots, X_{t_n} on $(\mathbb{R}^d)^n$, for any $n \geq 1$, and $t_1, \dots, t_n \in T$.

Proof:

For $\xi = (\lambda_1, \dots, \lambda_n) \in (\mathbb{R}^d)^n$, we set

$$\begin{aligned} (1.14) \quad \varphi(\xi) &\stackrel{\text{def}}{=} E\left[\exp\left\{i \sum_{j=1}^n \lambda_j \cdot X_{t_j}\right\}\right] = \exp\left\{-\frac{1}{2} E\left[\left(\sum_{j=1}^n \lambda_j \cdot X_{t_j}\right)^2\right]\right\} \\ &\quad \leftarrow \text{real valued centered Gaussian variable} \\ &= \exp\left\{-\frac{1}{2} \sum_{i, j=1}^n E\left[(\lambda_i \cdot X_{t_i}) (\lambda_j \cdot X_{t_j})\right]\right\} \\ &\quad \quad \quad \text{row vector} \quad \quad \quad \text{column vector} \\ &= \exp\left\{-\frac{1}{2} \sum_{i, j=1}^n \lambda_i \cdot E\left[X_{t_i} \cdot X_{t_j}\right] \lambda_j\right\} \stackrel{(1.12)}{=} \exp\left\{-\frac{1}{2} \sum_{i, j=1}^n \lambda_i \cdot \Gamma(t_i, t_j) \lambda_j\right\} \\ &\quad \quad \quad \text{d} \times \text{d matrix} \end{aligned}$$

But the characteristic function $\varphi(\cdot)$ completely determines the law of $(X_{t_1}, \dots, X_{t_n})$ on $(\mathbb{R}^d)^n$. □

We will now provide a characterization of Brownian motion as a continuous centered Gaussian process.

Proposition:

Let $B_t, t \geq 0$, be an \mathbb{R}^d -valued process defined on (Ω, \mathcal{A}, P) with P.a.s. continuous trajectories.

(1.15) $B_t, t \geq 0$, is a Brownian motion $\Leftrightarrow B_t, t \geq 0$ is a centered Gaussian process with $\Gamma(s, t) = (\Delta t) \overset{\uparrow}{\text{identity matrix}} \mathbb{I}_{d \times d}$

Proof:

$$\Rightarrow: \text{for } 0 \leq t_1 < t_2 < \dots < t_n, \lambda_1, \dots, \lambda_n \in \mathbb{R}^d, \text{ P.a.s.},$$

$$a \stackrel{\text{def}}{=} \sum_{i=1}^n \lambda_i \cdot B_{t_i} \stackrel{(1.11)}{=} \sum_{i=1}^n \lambda_i \cdot \left(\sum_{j=1}^i B_{t_j} - B_{t_{j-1}} \right) = \sum_{j=1}^n \left(\sum_{i=j}^n \lambda_i \right) \cdot (B_{t_j} - B_{t_{j-1}}).$$

By (1.1), the $(B_{t_j} - B_{t_{j-1}}), 1 \leq j \leq n$, are independent respectively $N(0, (t_j - t_{j-1}) \mathbb{I}_{d \times d})$ -distributed. Therefore a is a real valued centered Gaussian variable (use characteristic function), we have shown that $B_t, t \geq 0$, is a centered Gaussian process. Moreover for $0 \leq s \leq t$, we have

$$(1.16) \Gamma(s, t) = E[B_s^t B_t^t] = E[B_s^t B_s^t] + E[B_s^t (B_t - B_s)^t]$$

$$= \Delta \mathbb{I}_{d \times d} = (\Delta t) \mathbb{I}_{d \times d}.$$

$\underbrace{\quad}_{=0} \quad \leftarrow \text{independent and centered}$

Then for $t \leq s$, $\Gamma(s, t) = {}^t \Gamma(s, t) \stackrel{(1.16)}{=} (t \Delta) \mathbb{I}_{d \times d} = (\Delta t) \mathbb{I}_{d \times d}$.

\Leftarrow : If $0 \leq t_1 < \dots < t_n$ are given and in some auxiliary probability space $Y_j, 1 \leq j \leq n$, are independent $N(0, (t_j - t_{j-1}) \mathbb{I}_{d \times d})$ -distributed, we can define $X_j = \sum_{k=1}^j Y_k$.

The argument above shows that $X_j, 1 \leq j \leq n$, is a centered Gaussian process, and for $1 \leq i < j \leq n$:

$$(1.17) \Gamma(i, j) \stackrel{\text{def}}{=} E[X_i^t X_j^t] = E[X_i^t X_i^t] + E[X_i^t (X_j - X_i)^t]$$

$$= E\left[\left(\sum_{k=1}^i Y_k\right)^t \left(\sum_{k=1}^j Y_k\right)^t\right] \stackrel{\substack{\text{indep.} \\ \text{centered}}}{=} \sum_{1 \leq k \leq i} E[Y_k^t Y_k^t] = t_i \mathbb{I}_{d \times d}.$$

$\underbrace{\quad}_{\text{independent and centered}}$

As before (1.16), we thus see that:

$$(1.18) \Gamma(i, j) = (t_i \wedge t_j) I_{d \times d}, \text{ for } 1 \leq i, j \leq n.$$

By (1.13) we thus see that (X_1, \dots, X_n) has same law as $(B_{t_1}, \dots, B_{t_n})$. Therefore $(B_{t_1}, B_{t_2} - B_{t_1}, \dots, B_{t_n} - B_{t_{n-1}})$ has same law as $(X_1, X_2 - X_1, \dots, X_n - X_{n-1}) = (Y_1, \dots, Y_n)$. Hence (1.1) ii), iii) follow. Moreover (1.1) iv) holds by assumption and $E[B_0^t B_0^t] = 0$ implies by looking at the diagonal coefficients of this matrix that P.-a.s., $B_0 = 0$. This proves that $B_t, t \geq 0$, is a Brownian motion. \square

The above characterization is very helpful.

Examples:

1) Invariance by scaling:

Consider $B_t, t \geq 0$, an \mathbb{R}^d -valued Brownian motion, $\lambda > 0$, then (1.19) $\lambda B_{t/\lambda^2}, t \geq 0$, is also an \mathbb{R}^d -valued Brownian motion.

Indeed:

$B_t^\lambda \stackrel{\text{def}}{=} \lambda B_{t/\lambda^2}, t \geq 0$, is also a continuous centered gaussian process and for $\Delta \geq 0, t \geq 0$:

$$E[B_\Delta^\lambda {}^t B_t^\lambda] = \lambda^2 E[B_{\Delta/\lambda^2} {}^t B_{t/\lambda^2}] = \lambda^2 \begin{pmatrix} \Delta \wedge t \\ \lambda^2 & \lambda^2 \end{pmatrix} I_{d \times d} = (\Delta \wedge t) I_{d \times d},$$

and (1.19) follows from (1.15).

2) Invariance by time inversion:

Consider $B_t, t \geq 0$, an \mathbb{R}^d -valued Brownian motion and define

$$(1.20) \beta_s = 0, \text{ and } \beta_s = \Delta B_{1/s}, \text{ for } s > 0.$$

Then one has:

$$(1.21) \beta_s, s > 0, \text{ is an } \mathbb{R}^d\text{-valued Brownian motion.}$$

Indeed

$\beta_s, s > 0$, is a centered gaussian process, and for $0 < s, t$:

$$(1.22) E[\beta_s {}^t \beta_t] = \Delta t E[B_{1/s} {}^t B_{1/t}] = \Delta t \begin{pmatrix} 1 \wedge 1 \\ \Delta & \Delta \end{pmatrix} I_{d \times d} = \frac{\Delta t}{\Delta s t} I_{d \times d} = (\Delta \wedge t) I_{d \times d},$$

and this formula immediately extends to the case $0 \leq s, t$.

There only remains to see that P-a.s., $s \geq 0 \rightarrow \beta_s$ is continuous to conclude (1.21). To this end we note that by (1.22) and (1.13),
 (1.23) The laws of $\beta_s, s \geq 0$, and $B_t, t \geq 0$, on $C([0, \infty), \mathbb{R}^d)$ are identical. ↑
open in 0!

We let $X_u, u \geq 0$, denote the canonical process on $C([0, \infty), \mathbb{R}^d)$, and $\mathcal{G} = \sigma(X_u, u \geq 0)$ the canonical σ -algebra. If Q stands for the common law on $(C([0, \infty), \mathbb{R}^d), \mathcal{G})$ of $\beta_s, s \geq 0$, or $B_t, t \geq 0$, then
 (1.24) $\{ \lim_{u \rightarrow 0} X_u = 0 \} \in \mathcal{G}$ (it is an "event").

Indeed one has:

$$\{ \lim_{u \rightarrow 0} X_u = 0 \} = \bigcap_{n \geq 1} \bigcup_{m \geq 1} \bigcap_{u \in \mathbb{Q} \cap (0, \frac{1}{m})} \{ |X_u| \leq \frac{1}{n} \} \in \mathcal{G}.$$

As a result we find

$$Q \left(\lim_{u \rightarrow 0} X_u = 0 \right) = P \left[\lim_{u \rightarrow 0} B_u = 0 \right] \stackrel{(1.1) \text{ i), iv)}}{=} 1$$

|| (1.23)

$P \left[\lim_{u \rightarrow 0} B_u = 0 \right]$, and hence $\beta_s, s \geq 0$, fulfills (1.1) iv) as well. \square

Construction of Brownian motion:

We are now going to construct a Brownian motion on some (Ω, \mathcal{A}, P) . It suffices to consider the case $d=1$, since by taking d independent copies of a real valued Brownian motion, we obtain a d -dimensional Brownian motion.

We follow a method of Paul Lévy (48), later simplified by Z. Ciesielski (61).

We recall the Haar functions on \mathbb{R}_+ :

$$(1.25) \quad \varphi_\ell(t) = 1_{[\ell, \ell+1)}(t), \quad \ell \in \mathbb{N},$$

$$\varphi_{m,k}(t) = 2^{\frac{m}{2}} 1_{\left[\frac{k}{2^m}, \frac{k+1}{2^m}\right)}(t) - 2^{\frac{m}{2}} 1_{\left[\frac{k}{2^m} + \frac{1}{2^{m+1}}, \frac{k+1}{2^m}\right)}(t),$$

with $m, k \in \mathbb{N}$.

Fact:

(1.26) The $\varphi_\ell, \ell \geq 0, \varphi_{m,k}, m, k \geq 0$, form a complete orthonormal basis of $L^2(\mathbb{R}_+, dt)$.

Indeed:

- The functions $\varphi_\ell, \varphi_{m,k}$ have unit $L^2(\mathbb{R}_+, dt)$ -norms
- They are pairwise orthogonal in $L^2(\mathbb{R}_+, dt)$.
- The L^2 -closure of the span of the $\varphi_\ell, \varphi_{m,k}$ is $L^2(\mathbb{R}_+, dt)$, because one sees by induction on $m \geq 0$, that all $1_{\left[\frac{j}{2^m}, \frac{j+1}{2^m}\right)}, j \geq 0$, belong to the space generated by $\varphi_\ell, \ell \geq 0$, and $\varphi_{m',k}, 0 \leq m' < m, 0 \leq k$, and the above claim follows.

Heuristic (non-rigorous) description of the construction of B.M.

The idea is to use the formal development of \dot{B} (the derivative of Brownian motion!!!) in the above Haar basis. Formally we have:

$$\dot{B}(\cdot) \stackrel{''}{=} \sum_{\ell \geq 0} \varphi_\ell(\cdot) \left(\int_0^\infty \dot{B} \varphi_\ell dt \right) + \sum_{m,k \geq 0} \varphi_{m,k}(\cdot) \left(\int_0^\infty \dot{B} \varphi_{m,k} dt \right).$$

We then write

$$(1.27) \quad B(t) \stackrel{''}{=} \int_0^t \dot{B}(u) du \stackrel{''}{=} \sum_{\ell \geq 0} \int_0^t \varphi_\ell(u) du \left(\int_0^\infty \dot{B} \varphi_\ell dt \right) + \sum_{m,k \geq 0} \int_0^t \varphi_{m,k}(u) du \left(\int_0^\infty \dot{B} \varphi_{m,k} dt \right).$$

We now define

$$(1.28) \quad \zeta_\ell \stackrel{''}{=} \int_0^\infty \dot{B} \varphi_\ell dt \stackrel{''}{=} \int_0^{\ell+1} \dot{B} dt = B(\ell+1) - B(\ell)$$

$$(1.29) \quad \zeta_{m,k} \stackrel{''}{=} \int_0^\infty \dot{B} \varphi_{m,k} dt \stackrel{''}{=} 2^{\frac{m}{2}} \left(\int_{\frac{k}{2^m}}^{\frac{k+1}{2^m}} \dot{B} dt - \int_{\frac{k}{2^m} + \frac{1}{2^{m+1}}}^{\frac{k+1}{2^m}} \dot{B} dt \right)$$

$$\stackrel{''}{=} 2^{\frac{m}{2}} \left(B\left(\frac{k+1}{2^m}\right) - B\left(\frac{k}{2^m}\right) \right) - 2^{\frac{m}{2}} \left(B\left(\frac{k+1}{2^m}\right) - B\left(\frac{k}{2^m} + \frac{1}{2^{m+1}}\right) \right).$$

Now if a Brownian motion exists, then the right-hand sides of (1.28) and (1.29) make sense and the above $\zeta_\ell, \zeta_{m,k}$ are $N(0,1)$ -variables, and form a centered Gaussian family, (since they are linear combinations of $B(t), t \geq 0$). Moreover the variables $\zeta_\ell, \ell \geq 0, \zeta_{m,k}, m, k \geq 0$, are pairwise orthogonal:

$$E[\zeta_l \zeta_{l'}] = 0, l \neq l', \quad E[\zeta_l \zeta_{m,k}] = 0, \text{ for } l \neq l' \geq 0, m, k \geq 0,$$

$$E[\zeta_{m,k} \zeta_{m',k'}] = 0, \text{ for } (m,k) \neq (m',k').$$

This orthogonality follows from (1.1) ii), iii), (for instance $E[\zeta_l \zeta_{l'}] = 0$, for $l \neq l'$ is immediate to check, likewise $E[\zeta_l \zeta_{m,k}] = 0$, if $[\frac{k}{2^m}, \frac{k+1}{2^m}) \cap [l, l+1) = \emptyset$ is immediate, and on the other hand when the intersection is not empty, then $[\frac{k}{2^m}, \frac{k+1}{2^m}) \subseteq [l, l+1)$, and one has

$$E[\zeta_l \zeta_{m,k}] = 2^{\frac{m}{2}} E[(B(l+1) - B(l)) \left\{ \frac{B(k+1) - B(\frac{k}{2})}{2^m} - \frac{B(\frac{k}{2}) - B(\frac{k}{4})}{2^m} \right\} \left\{ \frac{B(\frac{k}{2}) - B(\frac{k}{4})}{2^m} - \frac{B(\frac{k}{4}) - B(\frac{k}{8})}{2^m} \right\} \right]$$

and writing

$$B(l+1) - B(l) = \underbrace{B(l+1) - B(\frac{l+1}{2})}_{2^m} + \underbrace{B(\frac{l+1}{2}) - B(\frac{l}{2})}_{2^m} + \underbrace{B(\frac{l+1}{2}) - B(\frac{l}{2})}_{2^m} - \underbrace{B(\frac{l}{2})}_{2^m} + \underbrace{B(\frac{l}{2}) - B(l)}_{2^m}$$

we can use (1.1) ii), iii) to conclude that $E[\zeta_l \zeta_{m,k}] = 0$ as well, the last equality " $E[\zeta_{m,k} \zeta_{m',k'}] = 0$ for $(m,k) \neq (m',k')$ is shown with analogous considerations).

Since the $\zeta_l, \zeta_{m,k}$ form a centered Gaussian family, are $N(0,1)$ -distributed and are pairwise uncorrelated, they are in fact $N(0,1)$ -distributed, (cf. (1.13)). Hence the formal formulas (1.27), (1.28), (1.29) tell us where we should "look for a Brownian motion".

We will now see how the above non-rigorous considerations can be transformed into a real proof.

Mathematical construction.

We consider on some suitable probability space (Ω, \mathcal{Q}, P) a (countable) family of iid $N(0,1)$ -distributed variables $\zeta_l, \zeta_{m,k}$, $l \geq 0, m, k \geq 0$, (for instance $\Omega = (0,1)$, $\mathcal{Q} = \mathcal{B}((0,1))$, $P =$ Lebesgue measure on $(0,1)$, will do the job).

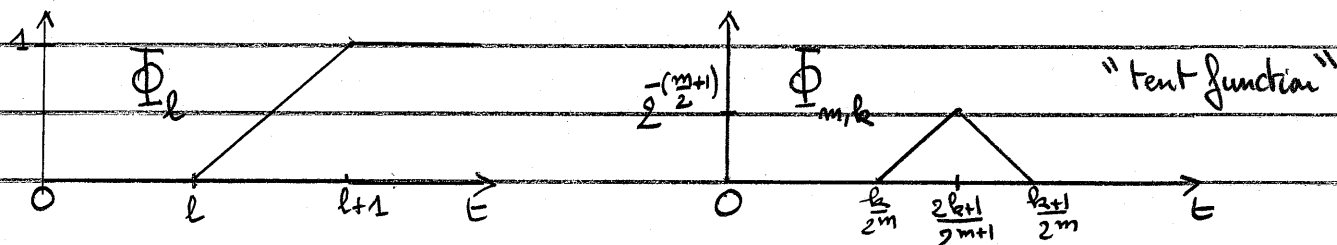
We then define for $n \geq 1$ and $t \geq 0$:

$$(1.30) X_n(t) = \sum_{0 \leq l < n} \Phi_l(t) \zeta_l + \sum_{0 \leq m < n} \left(\sum_{0 \leq k < n 2^m} \Phi_{m,k}(t) \zeta_{m,k} \right),$$

where

$$(1.31) \Phi_l(t) = \int_0^t \varphi_l(u) du \quad \text{and} \quad \Phi_{m,k}(t) = \int_0^t \varphi_{m,k}(u) du$$

(these are called Schauder functions)



Lemma:

(1.32) P-a.s., $X_n(\cdot, \omega)$ converges uniformly on compact intervals of \mathbb{R}_+ to a finite limit $X(\cdot, \omega)$.

Proof:

It suffices to prove that for any $n_0 \geq 1$, P-a.s., $X_n(\cdot, \omega)$ converges uniformly on $[0, n_0]$ to a finite limit, (because we can then find $N \in \mathcal{A}$, with $P[N] = 0$, so that on N^c , $X_n(\cdot, \omega)$ converges uniformly on any $[0, n_0]$, and then define

$$X(t, \omega) = \lim_n X_n(t, \omega), \quad \omega \in N^c, t \geq 0, \\ = 0, \quad \text{when } \omega \in N, t \geq 0).$$

For $t \in [0, n_0]$, we have for $n \geq n_0$:

$$(1.33) X_n(t) = \sum_{l=0}^{n_0-1} \Phi_l(t) \zeta_l + \sum_{0 \leq m < n} \left(\sum_{0 \leq k < n 2^m} \Phi_{m,k}(t) \zeta_{m,k} \right),$$

and for each $t \in [0, n_0]$, and each $m \geq 0$,

there is at most one $k \geq 0$, such that $\Phi_{m,k}(t) \neq 0$.

We then define:

$$(1.34) a_m(\omega) = \sup_{t \in [0, n_0]} \left| \sum_{k < n_0 2^m} \Phi_{m,k}(t) \zeta_{m,k} \right| \leq 2^{-(\frac{m+1}{2})} \sup_{k < n_0 2^m} |\zeta_{m,k}|.$$

We will control this supremum. To this end we note that for $\zeta \sim N(0,1)$ -distributed

$$(1.35) P[|\zeta| > a] \leq \sqrt{\frac{2}{\pi}} \frac{1}{a} \exp\left\{-\frac{a^2}{2}\right\}, \quad \text{for } a > 0,$$

$$\text{(indeed } \int_0^\infty \exp\left\{-\frac{(a+u)^2}{2}\right\} du \leq \int_0^\infty \exp\left\{-\frac{a^2}{2} - au\right\} du = \frac{1}{a} \exp\left\{-\frac{a^2}{2}\right\} \text{)}.$$

It follows that

$$\sum_{m \geq 1} P \left[\sup_{0 \leq k < n_0 2^m} |\zeta_{m,k}| > \sqrt{2m} \right] \leq \sum_{m \geq 1} \sqrt{\frac{2}{\pi}} \frac{n_0 2^m}{\sqrt{2m}} e^{-m} < \infty.$$

Thus Borel-Cantelli's lemma implies that for P-a.e. ω , there exists $m_0(\omega)$ such that for $m \geq m_0(\omega)$, $\sup_{k \leq n_0 2^m} |\zeta_{m,k}(\omega)| \leq \sqrt{2m}$.

As a result:

$$(1.36) \text{ P-a.s. } \sum_{m \geq m_0(\omega)} \frac{1}{m} a_m(\omega) \leq \sum_{m \geq m_0(\omega)} 2^{-(\frac{m+1}{2})} \sqrt{2m} < \infty.$$

It follows that P-a.s., $X_n(\cdot, \omega)$ converges uniformly on $[0, n_0]$ to a finite limit. \square

We will now see that the above defined $X(t, \omega)$, $t \geq 0$, is a Brownian motion. First we observe that each $X_m(t)$, $t \geq 0$, is a centered Gaussian process, (the $X_m(t)$ are finite linear combinations of the iid $N(0,1)$ -distributed $\zeta_l, l \geq 0, \zeta_{m,k}, m, k \geq 0$).

Note also that

$$(1.37) \text{ for } t \geq 0, X_m(t, \omega) \xrightarrow{L^2(\Omega, \mathcal{G}, P)} X(t, \omega).$$

Indeed the $\zeta_l, \zeta_{m,k}$ are orthogonal in $L^2(P)$ so that, cf (1.30), $X_m(t)$ and $X_{m+m}(t) - X_m(t)$ are orthogonal, and we only need to check that $\sup_n E[X_n^2(t)] < \infty$, that is:

$$(1.38) \sum_{l \geq 0} \Phi_l(t)^2 + \sum_{m, k \geq 0} \Phi_{m,k}(t)^2 < \infty.$$

To check this last point we observe that the above sum equals:

$$(1.39) \sum_{l \geq 0} \left(\int_0^t \varphi_l(u) du \right)^2 + \sum_{m, k \geq 0} \left(\int_0^t \varphi_{m,k}(u) du \right)^2 \stackrel{(1.26)}{=} \text{Parseval relation}$$

$$\|1_{E[0,t]}\|_{L^2(\mathbb{R}_+, du)}^2 = t < \infty.$$

In a very similar vein we calculate $E[X_m(s)X_m(t)]$ for $0 \leq s, t$, as follows:

$$(1.40) \quad E[X_n(\Delta) X_n(t)] \stackrel{(1.30)}{=} \sum_{0 \leq l < m} \frac{\Phi(t) \Phi(\Delta)}{l} + \sum_{\substack{0 \leq m < n \\ 0 \leq k < n 2^m}} \frac{\Phi(t) \Phi(\Delta)}{m, k} \frac{\Phi(\Delta)}{m, k}$$

$$= \langle \pi_n(1_{[0, t]}) , \pi_n(1_{[0, \Delta]}) \rangle_{L^2(\mathbb{R}_+, du)}$$

with π_n the orthogonal projection in $L^2(\mathbb{R}_+, du)$ on the space spanned by $\varphi_l, 0 \leq l < n, \varphi_{m, k}, 0 \leq m < n, 0 \leq k < n 2^m$.

Combining (1.37) and (1.40), we find that

$$E[X_n(\Delta) X_n(t)] = \langle \pi_n(1_{[0, \Delta]}) , \pi_n(1_{[0, t]}) \rangle_{L^2(\mathbb{R}_+, du)}$$

$$\downarrow n \rightarrow \infty \qquad \qquad \qquad \downarrow n \rightarrow \infty$$

$$E[X(\Delta) X(t)] = \langle 1_{[0, \Delta]} , 1_{[0, t]} \rangle_{L^2(\mathbb{R}_+, du)} = \Delta t.$$

Note that weak limits of centered Gaussian distributions are centered Gaussian, (take characteristic functions). It follows that linear combinations of $X(t)$, which are limit, in $L^2(P)$ by (1.37), and thus in distribution, of linear combinations of $X_n(t)$, are centered Gaussian variables.

So $X(t, \omega), t \geq 0$, is a centered Gaussian process. From the above calculation $\Gamma(s, t) = \Delta \min(s, t)$, and due to (1.32) P-a.s., $t \geq 0 \rightarrow X(t, \omega)$ is continuous.

We have thus proved that $X(t, \omega)$ is a Brownian motion on the probability space (Ω, \mathcal{A}, P) selected at the bottom of p. 14.