

Lecture 19: (Chap 6, cont.)

Complement

We will now present Novikov's criterion, which refines the condition we gave in (6.29) to ensure that Z_t is a martingale and not merely a continuous local martingale.

Theorem: (Novikov's criterion)

Let $M_t, t \geq 0$, be a continuous local martingale with $M_0 = 0$, such that

$$(6.72) \quad E \left[\exp \left\{ \frac{1}{2} \langle M \rangle_\infty \right\} \right] < \infty,$$

then

$$(6.73) \quad E \left[\exp \left\{ \frac{1}{2} \sup_{t \geq 0} |M_t| \right\} \right] < \infty,$$

and

$$(6.74) \quad Z_t = \exp \left\{ M_t - \frac{1}{2} \langle M \rangle_t \right\}, t \geq 0, \text{ is a uniformly integrable continuous martingale,}$$

(and of course $M_t, t \geq 0$, is a continuous martingale as well).

Remark:

If instead of (6.72) we assume that for some $T > 0$,

$$(6.75) \quad E \left[\exp \left\{ \frac{1}{2} \langle M \rangle_T \right\} \right] < \infty,$$

the above theorem can be applied to $M_{t \wedge T}, t \geq 0$, and we find that

$$(6.76) \quad Z_t = \exp \left\{ \frac{1}{2} M_t - \frac{1}{2} \langle M \rangle_t \right\}, t \leq T, \text{ is a continuous martingale. } \square$$

Proof:

We first observe that $E[\langle M \rangle_\infty] < \infty$ implies that

$$(6.77) \quad M_t, t \geq 0, \text{ is a continuous martingale bounded in } L^2.$$

Indeed (this is just as in the exercise below (6.30)), one chooses a sequence $T_n \uparrow \infty$ of stopping times so that $M_{t \wedge T_n}, t \geq 0$, are bounded martingales. Then one has

$$E[M_{t \wedge T_n}^2] \stackrel{(5.36)}{=} E[\langle M \rangle_{t \wedge T_n}] \leq E[\langle M \rangle_\infty], \text{ and by Fatou's lemma}$$

$$E[M_t^2] \leq \liminf_n E[M_{t \wedge T_n}^2] \leq E[\langle M \rangle_\infty].$$

It now also follows with similar considerations as in (4.104) and Doob's inequality that $M_t, t \geq 0$, is a continuous martingale, with $E[\sup_{t \geq 0} |M_t|^2] < \infty$, and that $M_\infty = \lim_{t \rightarrow \infty} M_t$ is well-defined, by the martingale convergence theorem

(6.73). Note that for $0 \leq t < \infty$,

$$(6.78) \quad E[\exp\{\frac{1}{2} M_t\}] = E[\exp\{\frac{1}{2} M_t - \frac{1}{4} \langle M \rangle_t\} \cdot \exp\{\frac{1}{4} \langle M \rangle_t\}] \\ \stackrel{\text{Cauchy-Schwarz}}{\leq} E[\exp\{M_t - \frac{1}{2} \langle M \rangle_t\}]^{1/2} E[\exp\{\frac{1}{2} \langle M \rangle_t\}]^{1/2}.$$

Since $Z_t, t \geq 0$, is a non-negative local martingale, it is also a supermartingale, (see exercise below (6.67): for $0 \leq s < t$

$$E[Z_t | \mathcal{G}_s] = E[\lim_n Z_{t \wedge T_n} | \mathcal{G}_s] \stackrel{\text{Fatou}}{\leq} \liminf_n E[Z_{t \wedge T_n} | \mathcal{G}_s] \\ = \liminf_n Z_{s \wedge T_n} = Z_s, \text{ thus proving the claim.}$$

As a result $E[Z_t] \leq E[Z_0] = 1$, and coming back to (6.78)

$$(6.79) \quad E[\exp\{\frac{1}{2} M_t\}] \leq E[\exp\{\frac{1}{2} \langle M \rangle_t\}]^{1/2}, \quad 0 \leq t < \infty.$$

The same argument applied to $-M$ yields that

$$\sup_{t \geq 0} E[\cosh(\frac{1}{2} M_t)] \leq E[\exp\{\frac{1}{2} \langle M \rangle_\infty\}]^{1/2}, \text{ so that} \\ (6.80) \quad E[\cosh(c M_t)] \leq E[\exp\{\frac{1}{2} \langle M \rangle_\infty\}]^{1/2}, \text{ for } 0 \leq c \leq \frac{1}{2}, 0 \leq t < \infty, \\ \text{and in particular since } \cosh x \geq \frac{1}{2} e^x,$$

$$(6.81) \quad \sup_{0 \leq t < \infty} E[e^{c M_t}] < \infty, \text{ for } 0 \leq c \leq \frac{1}{2}.$$

Jensen's inequality implies that $e^{c M_t}, t \geq 0$, is a non-negative submartingale. It then follows from Doob's inequality (4.76)

with $p = \frac{1}{2c}$, for $0 < c < \frac{1}{2}$, that

$$(6.82) \quad E[\sup_{t \geq 0} \exp\{\frac{1}{2} M_t\}] = E[(\sup_{t \geq 0} \exp\{c M_t\})^p] \\ \stackrel{(4.76)}{\leq} \left(\frac{p}{p-1}\right)^p \sup_{t \geq 0} E[\exp\{p c M_t\}]$$

$$= \left(\frac{p}{p-1}\right)^p \sup_{t \geq 0} E[\exp\{\frac{1}{2} M_t\}] \stackrel{(6.81)}{<} \infty.$$

Of course a similar bound holds for $-M$ in place of M . Note also

$$\text{that } \sup_{t \geq 0} (\cosh(\frac{1}{2} M_t)) \leq \frac{1}{2} \sup_{t \geq 0} e^{\frac{1}{2} M_t} + \frac{1}{2} \sup_{t \geq 0} e^{-\frac{1}{2} M_t},$$

and hence

$$(6.83) \quad E[\sup_{t \geq 0} \cosh(\frac{1}{2} M_t)] < \infty,$$

from which follows that $E[\exp\{\frac{1}{2} \sup_{t \geq 0} |M_t|\}] < \infty$, and (6.73) holds

(6.74):

We will use the next

Lemma:

(6.84) If $E[Z_\infty] = 1$, then $Z_t, 0 \leq t < \infty$, is a uniformly integrable martingale.

Proof:

From the supermartingale property of $Z_t, t \geq 0$, follows that $1 = E[Z_\infty] \leq E[Z_t] \leq E[Z_0]$, for $0 \leq t < \infty$, so that $E[Z_t] = 1$, for $0 \leq t < \infty$.

Note that P.-a.s., $Z_{t \wedge T_n} \xrightarrow{n \rightarrow \infty} Z_t$, and $E[Z_{t \wedge T_n}] = E[Z_t] = 1$, and these variables are non negative.

It now follows that they are uniformly integrable and

$$(6.85) \quad Z_{t \wedge T_n} \xrightarrow[n \rightarrow \infty]{L^1} Z_t, \quad \text{for } 0 \leq t < \infty,$$

see for instance Durrett "Probability: Theory and Examples", p. 224.

This now implies that $Z_t, t \geq 0$, is a martingale. Moreover

since P.-a.s., $Z_t \rightarrow Z_\infty$, as $t \rightarrow \infty$, and $E[Z_t] = E[Z_\infty] = 1$,

the above argument shows that

$$(6.86) \quad Z_t \xrightarrow{L^1} Z_\infty, \quad \text{as } t \rightarrow \infty,$$

so that

$$(6.87) \quad Z_t = E[Z_\infty | \mathcal{G}_t], \quad \text{for } t \geq 0,$$

and the conclusion of the lemma follows, cf. Durrett, p. 223. \square

We will now show that

$$(6.88) \quad E[Z_\infty] \geq 1,$$

since we already know that $E[Z_\infty] \leq 1$, the claim (6.74) will now follow from the lemma.

By (6.29), (6.30), with $T = \infty$, we know from (6.72) that

$$E[\exp\{a M_\infty - \frac{a^2}{2} \langle M \rangle_\infty\}] = 1, \quad \text{for } a \in [0, 1].$$

Note that the following equality holds:

$$\exp\left\{a M_\infty - \frac{a^2}{2} \langle M \rangle_\infty\right\} = \exp\left\{M_\infty - \frac{1}{2} \langle M \rangle_\infty\right\}^{a^2} \exp\left\{\frac{a}{1+a} M_\infty\right\}^{1-a^2}.$$

Using Hölder's inequality with $p = a^{-2}$ and $q = (1-a^2)^{-1}$, we find:

$$(6.89) \quad 1 \leq E[Z_\infty]^{a^2} E\left[\exp\left\{\frac{a}{1+a} M_\infty\right\}\right]^{1-a^2}.$$

Using (6.73), we can use dominated convergence to argue that

$$\lim_{a \rightarrow 1} E\left[\exp\left\{\frac{a}{1+a} M_\infty\right\}\right] = E\left[\exp\left\{\frac{1}{2} M_\infty\right\}\right] \in (0, \infty).$$

As a result we obtain that

$$(6.90) \quad \lim_{a \rightarrow 1} E\left[\exp\left\{\frac{a}{1+a} M_\infty\right\}\right]^{1-a^2} = 1,$$

and the claim (6.88) now follows from (6.89) and (6.90).

This concludes the proof of (6.74). \square

CHAPTER 7: STOCHASTIC DIFFERENTIAL EQUATIONS AND MARTINGALE PROBLEMS

We begin with some heuristic considerations.

In this chapter we want to construct processes, which "locally near a point x in \mathbb{R}^d ", move like the process

$$x + t b(x) + \sigma(x) B_t,$$

where $b(x) \in \mathbb{R}^d$, $\sigma(x)$ is a $(d \times n)$ -matrix and B_t is an n -dimensional Brownian motion. We want that "infinitesimally" the increment of the process we construct behaves as a Gaussian variable with mean: $b(x)dt$, and

covariance matrix: $a(x)dt$,

where for $\xi \in \mathbb{R}^d$

$$\begin{aligned} \xi^t a(x) \xi dt &= E \left[\left(\xi^t \sigma(x) \cdot (B_{t+dt} - B_t) \right)^2 \right] \\ &= E \left[\left(\xi^t \underbrace{\sigma(x)}_{\substack{\text{n-vector} \\ \text{matrix}}} \cdot (B_{t+dt} - B_t) \right)^2 \right] = \xi^t \sigma(x)^t \sigma(x) \xi dt \\ &= \xi^t \sigma(x)^t \sigma(x) \xi dt, \end{aligned}$$

i.e. in other words $a(x) = \sigma(x)^t \sigma(x)$ ($d \times d$ -matrix).

We will consider two approaches to build such processes. The first approach will rely on solving stochastic differential equations (SDE):

$$X_t^i = x_0^i + \int_0^t b_i(X_s) ds + \sum_{j=1}^n \int_0^t \sigma_{ij}(X_s) dB_s^j, \quad i=1, \dots, d,$$

or in vector notation:

$$(7.1) \quad X_t = x + \int_0^t b(X_s) ds + \int_0^t \sigma(X_s) \cdot dB_s, \quad x \in \mathbb{R}^d.$$

The second approach will be based on a martingale problem,

i.e. finding on $(C(\mathbb{R}_+, \mathbb{R}^d), \mathcal{F}_t)$ a probability P_x , such that

$$(7.2) \quad \left\{ \begin{array}{l} M_t^f \stackrel{\text{def}}{=} f(X_t) - f(X_0) - \int_0^t L f(X_s) ds, \quad t \geq 0, \text{ is an } (\mathcal{F}_t)\text{-martingale under } P_x, \\ \text{for all } f \in C_c^2(\mathbb{R}^d), \text{ with } L f(y) \stackrel{\text{def}}{=} \frac{1}{2} \sum_{i,j=1}^d a_{ij}(y) \frac{\partial^2}{\partial y_i \partial y_j} f(y) + \sum_{i=1}^d b_i(y) \frac{\partial}{\partial y_i} f(y), \quad y \in \mathbb{R}^d, \\ P_x[X_0 = x] = 1. \end{array} \right.$$

This latter approach shares the same spirit as Lévy's characterization of Brownian motion, cf. (6.44), (6.45).

Notation:

- For $b \in \mathbb{R}^d$, $|b| = \left(\sum_{i=1}^d b_i^2 \right)^{1/2}$,

- For $\sigma \in M_{d \times n}$, $|\sigma| = \left(\sum_{\substack{1 \leq i \leq d \\ 1 \leq j \leq n}} \sigma_{ij}^2 \right)^{1/2} = \left\{ \text{Trace} \left(\sigma \sigma^T \right) \right\}^{1/2} = \left\{ \text{Trace} \left(\sigma^T \sigma \right) \right\}^{1/2}$
 $\in M_{d \times d}$ $\in M_{n \times n}$

- $(\Omega, \mathcal{G}, (\mathcal{G}_t)_{t \geq 0}, \mathbb{P})$ a probability space satisfying the usual conditions, cf. (4.5), (4.6).

- $B_t, t \geq 0$, an n -dimensional (\mathcal{G}_t) -Brownian motion, or equivalently in view of (6.44), (6.45), for all $1 \leq i, j \leq n$,

$B_t^i, t \geq 0$, are continuous local martingales with $B_0 = 0$, and

$B_t^i B_t^j - \delta_{ij} t, t \geq 0$, are continuous local martingales.